

Online Appendix: Consumer Preferences for Cigarettes and Heated Tobacco Products in Japan: Evidence from a Discrete Choice Experiment

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In this Appendix we first check the robustness of the main results reported in the paper, and then discuss several lines of evidence that shed light on the validity of the stated preference (SP) data we collected through our discrete choice experiment (DCE), and the implications for the empirical results reported in the text of the paper.

In section A we estimate a mixed logit model of consumer product choices that allows for more flexible substitution pattern and conduct counterfactual prediction under different policy scenarios described in the paper as a robustness check for the main results estimated from conditional logit models reported in the paper. In section B we review previous research that compares SP and revealed preference (RP) data in a range of applications. In section C we review previous research that conducts DCEs of tobacco product choices. In Section D we provide additional empirical evidence on the validity of our SP data. Based on previous research and the empirical evidence in section D, we conclude that because tobacco products are familiar market goods, econometric models estimated using the SP data from our DCE are likely to provide reliable forecasts of consumer demand.

A. Robustness of Main Results

Table A1 provides estimated results from a mixed logit model of consumer tobacco product choices. The estimated distribution of consumer preferences exhibits a similar pattern to that estimated from the conditional logit model reported in the paper. Table A2 presents the predicted market shares of cigarettes, HTPs, and quitting under different policy scenarios considered in the paper, the results are similar to the predictions derived from the conditional logit model reported in the paper.

Table A1. Estimation results of a mixed logit model

	Immediate Choice Today		Choice of 6 Months from Now	
	Mean	SD	Mean	SD
Cigarettes	6.847***	4.584***	8.045***	5.686***

Alternative-Specific-Constant		(0.292)	(0.247)	(0.376)	(0.356)
Alternative-Specific-Constant	HTPs	5.668***	4.862***	6.944***	4.559***
		(0.312)	(0.293)	(0.353)	(0.260)
Price	Price in 100 JPY	-0.636***	0.670***	-0.854***	1.056**
		(0.047)	(0.107)	(0.067)	(0.187)
HTP available flavor	Tobacco, menthol, fruity/coffee/mint	0.261**	0.442**	0.350**	0.574***
		(0.125)	(0.211)	(0.137)	(0.208)
	Tobacco and menthol	0.072	0.352*	0.278**	0.133
		(0.124)	(0.213)	(0.133)	(0.268)
HTP available nicotine level	Up to 30mg	0.006	0.068	-0.035	0.905***
		(0.123)	(0.480)	(0.140)	(0.206)
	Up to 50mg	-0.259**	0.728***	-0.299**	0.675***
		(0.131)	(0.168)	(0.139)	(0.214)
HTP warning	Are not completely risk free	0.075	0.119	-0.059	0.568***
		(0.144)	(0.251)	(0.158)	(0.192)
	Contain nicotine, which is addictive	0.016	0.250	-0.315**	0.144
		(0.147)	(0.245)	(0.158)	(0.219)
	Smoking is a cause of lung cancer	0.146	0.735***	-0.027	0.069
		(0.148)	(0.219)	(0.153)	(0.274)
Log-likelihood at convergence			-3207		-3476
Number of subjects			523		523
Number of observations			18828		18828

Notes: The random coefficient on the negative of price is assumed to follow lognormal distribution; other random coefficients are assumed to follow normal distribution. The model is estimated in preference space by maximum simulated likelihood. 500 shifted and shuffled Halton draws are used for the simulation. Standard errors are in parentheses. Inference * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

Table A2. Predicted market shares under different counterfactual scenarios.

Counterfactual scenarios	Cigarettes	HTP	Quit
Status quo	0.486	0.398	0.116
Lower HTP tax	0.437	0.482	0.080
Higher HTP tax	0.499	0.375	0.126
HTP flavor ban	0.490	0.393	0.118
Higher HTP tax + HTP flavor ban	0.503	0.370	0.127

B. Literature Review of Studies that Compare SP and RP Data

DCEs and the related contingent valuation method are used to collect SP data in a range of applications. DCEs are commonly used in marketing research and economics to provide predictions of consumer demand

in scenarios that are not yet observed in actual markets. In addition to the tobacco product DCEs discussed below in section C, examples of the use of DCEs to study hypothetical market situations include studies of electricity markets (Blass, Lach, and Manski, 2010), health insurance markets (Kesternich, Hiess, McFadden, and Winter, 2013), labor markets (Mas, Alexandre and Pallais, 2017; Maestas et al. 2023), and firearms markets (Moshary, Shapiro, & Drango NBER Working Paper 2023). DCEs are also widely used in health economics to evaluate existing or prospective pharmaceutical products and health care treatment interventions (Ryan et al. 2007). Another large body of research uses DCEs and the related contingent valuation method to estimate willingness to pay for non-market goods like environmental quality.

Research that compares SP and RP data concludes that the external validity of SP data is much stronger in applications similar to familiar market goods. In a narrative review of DCE research, McFadden (2017) concludes that there is a “sharp reliability gradient”:

Forecasts that are comparable in accuracy to RP forecasts can be obtained from well-designed SP studies for familiar, relatively simple goods that are similar to market goods purchased by consumers, particularly when calibration to market benchmarks can be used to correct experimental distortions. However, studies of unfamiliar, complex goods give erratic, unreliable forecasts.

McFadden is therefore skeptical about SP data on complex and unfamiliar environmental public goods. Although he does not discuss health care applications, by the same reasoning DCEs might not provide reliable data on unfamiliar pharmaceutical and health care treatment interventions.

Penn and Hu (2018) report a meta-analysis that provides quantitative evidence consistent with McFadden’s (2017) conclusion that SP data are more reliable for familiar market goods. The meta-analysis used estimates from 132 studies that provided 908 observations of comparisons of SP and RP data. For studies including choice experiments that did not provide estimates of willingness to pay but did provide proportions of responses, Penn and Hu inferred lower-bound estimates of willingness to pay. Each observation is an estimate of the “calibration factor” (CF) which shows the ratio of willingness to pay estimated from SP data to the willingness to pay estimated from RP data. When SP and RP estimates are similar, the CF will be close to one. In the meta-analysis, about one quarter of the CFs are between 0.81 and 1.2. The distribution of CFs is skewed right showing a tendency for SP willingness-to-pay estimates to be larger than RP estimates, sometimes to a large extent. The median CF is 1.94, implying that for almost half of the observations the SP estimate is over twice as large as the RP estimate. Penn and Hu estimate regression models of the effects of study characteristics on CFs. The results imply that compared to studies of public goods, studies of private goods find lower CFs. The results also imply that compared to other hypothetical elicitation methods, CFs are lower for studies that used choice experiments. The meta-analysis empirical results are consistent with McFadden’s conclusion that although SP data from DCEs are reliable for private goods, there is a sharp reliability gradient for SP data on willingness to pay for public goods.

In the remainder of this section, we discuss examples of studies that compare SP and RP data on health-related choices. Quaife et al (2018) review a number of studies of health-related choices that focused on the external validity of DCEs (by comparing SP and RP estimates) and conclude that DCEs provide moderate levels of external validity in terms of matching actual choices. de Bekker-Grob et al (2020) find that when measured at the individual level, stated preferences in a DCE about vaccinations predict 91 percent of actual choices. Telser and Zweifel (2007) examine the external validity of a DCE focused on decisions about a harm reduction product (hip protectors for accidental falls). They compare the willingness to pay for risk reduction that was derived from the DCE to other measures of willingness to pay for the same risk reduction derived from established alternatives that used revealed preference data. The comparison supports a high level of convergent validity. Linley and Hughes (2013) examine hypothetical

decisions about new medicine approvals and find that the predicted probabilities of recommending new medicines derived from the DCE match well with the cumulative probability of actual positive recommendations (though the ability of the DCE to discriminate between individual new medicines was limited). Mahammad et al (2017) use a DCE and estimate a mixed logit based on the hypothetical choices with respect to type of tuberculosis treatment (or none) in response to each treatment having six treatment attributes. They compare these choices with actual choices and find strong external validity and the degree of accuracy depends on the distributional assumptions used in the mixed logit models with some models. Kesternich et al (2013) implement a DCE to analyze Medicare part D choices and compare these results to those that emerge from analysis of actual choices. They conclude that hypothetical choice experiments are useful in studying insurance choices as hypothetical behavior is related to actual behavior. They find that the coefficients that emerge in the DCE experiment are of the same sign as the coefficients that are estimated from market behavior. They note that the magnitudes of the coefficients are quite similar and do not find significant differences between hypothetical and real choices between different attributes of the insurance scenarios. They do find a higher willingness to pay for insurance in the hypothetical market and thus higher insurance take-up rates but attribute this to the nature of the default option in the DCE.

C. DCE Studies of Tobacco Product Choices

A growing body of research conducts DCEs to study the determinants of consumer choices about tobacco products. Table C1 lists recent DCE studies of tobacco product choices published in economics journals. Table C1 includes two studies that report results from exercises that use a combination of RP and SP data to develop a calibrated model that is grounded in real-world behavior. Kenkel, Peng, Pesko, and Wang (2020) report in an Online Appendix the results of a conditional logit model estimated using a combination of SP data from a DCE and RP data from the DCE subjects' responses about their prior use of e-cigarettes, combustible cigarettes, and nicotine replacement products. The estimated scale parameter is close to 1, suggesting that the scales in the RP and SP data are similar. The estimated coefficients on the tobacco product attributes show the same patterns as in the model based on SP data only.

Buckell and Hess (2019) report the results of a more in-depth investigation of combining SP and RP data on tobacco product choices. In a model estimated using combined SP and RP data they estimate a scale parameter greater than 1, consistent with the argument that in SP data subjects overstate the impact of interventions which leads to elasticities that are biased upwards. In terms of forecasts from the models, they find that compared to calibrated models the uncalibrated forecasts under-predict cigarette choices and over-predict e-cigarette choices. They conclude that appropriately calibrated choice models "provide better quality empirical evidence for policymakers." (Buckell and Hess 2019, p. 100)

The remaining tobacco product DCEs listed in Table C1 do not provide in-depth discussions of external validity, but the results of the studies are consistent with predictions from health economic models of consumer behavior. For example, Marti, Buckell, Maclean, and Sindelar (2019) conduct a DCE to study how smokers' product choices are affected by variations in the perceived healthiness and cessation effectiveness of e-cigarettes, as well as by bans on smoking in public places and prices. Buckell, Hensher, and Hess (2021) use SP data from a DCE combined with a latent variable approach to model addiction. They find that more addicted smokers are unwilling to switch to e-cigarettes. More broadly, the studies in Table C1 find that smokers' product choices respond to cigarette and e-cigarette prices; the estimated price-responsiveness is generally consistent with the large body of econometric estimates from observational (RP) data (DeCicca, Kenkel, and Lovenheim 2022).

In addition to the tobacco product DCEs listed in Table C1, Rousu, Marette, Thrasher, and Lusk (2014) report the results of non-hypothetical experimental auctions where smokers placed bids on cigarette

packages that carried either a text-only or graphic warning label. Depending on the location of the text-only warning label, 40 to 49 percent of the subjects bid more for the cigarette packages with text-only warning label than they bid for packages with the graphic warning label. The subjects' revealed lower willingness to pay for packages with graphic warning labels is consistent with our DCE results that graphic warning labels will reduce cigarette consumption. However, the experimental auction data do not include direct measures of consumer cigarette choices or their choices to attempt to quit smoking. The results are uninformative about whether the mechanisms through which graphic warning labels affected consumer willingness to pay. Rousu et al. (2014) propose a model that assumes the graphic warning labels provide information but acknowledge the limitation that "our model of the value of information does not make a distinction between different types of information effects and whether they are based on analytical reasoning or emotions."

DCE studies of tobacco product choices are also published in inter-disciplinary public health journals, including journals focusing on tobacco such as *Tobacco Control* and the *Journal of Nicotine and Tobacco Research*.¹ These articles are not included in our Table C1 review because they have a different focus for a different audience. For example, Shang, Huang, Chaloupka, and Emery (2018) focus on the roles flavors, e-cigarette device type, and e-cigarette warning labels play in youth stated preferences to try e-cigarettes. Subjects were presented with e-cigarette products with varying attributes but were not given the alternative of choosing combustible cigarettes.

In addition, we note that outside economics journals the term "discrete choice experiment" is used in both a broad and narrow sense. In the broad sense, DCE has been used to describe various surveys that asks subjects to make choices with random assignment of the descriptions of the alternatives. In contrast, Louviere, Flynn, and Carson (2010) define DCEs as being necessarily grounded in random utility theory. Some of the studies published in public health journals do not fit this narrow definition of DCEs, even though the studies use the term to describe their research method. For example, Reynolds, Popova, Ashley et al. (2022) report a DCE about very low nicotine cigarettes (VLNCs) that asked respondents which message would most motivate them and least motivate them to quit smoking; the message attributes varied in terms of content about VLNCs and the source of the message. Subjects did not make choices between products. This is a study of consumer perceptions of message effectiveness which cannot be grounded in random utility theory. A related concern is that in some DCEs where subjects make choices between tobacco products, the product attributes are described in terms of consumer perceptions rather than observable characteristics of the products and/or the product marketplace. For example, Shang, Weaver, White, et al. (2020) report a DCE about e-cigarettes that included "less harmful to health than cigarettes" and "effective for helping people quit" as product attributes. The Marti et al. (2019) study included in Table C1 also uses this approach. Although the results of these studies provide information about the relative importance of these perceptions on tobacco product choices, the results are less useful for policy analysis because perceptions are not directly policy manipulable. For this reason, in our DCE we describe the policy-manipulable attribute of warning labels.

Table C1. Recent DCE Studies of Tobacco Product Choices in Economics Journals

Authors	Year	Journal
Buckell, Hensher, and Hess	2021	Health Economics

¹ Regmi, Kaphle, Timilsina, and Tuha (2018) report a systematic review of peer-reviewed studies published from 2000 – 2016 that used DCE methods in tobacco control. Of the 12 studies included in their review, 4 were published in health economics journals. Because these 4 studies focused on pharmaceutical smoking cessation products and did not include e-cigarettes, they are not included in Table C1.

Buckell and Hess	2019	Journal of Health Economics
Kenkel, Peng, Pesko, and Wang	2020	Health Economics
Marti, Buckell, Maclean, and Sindelar	2019	Economic Inquiry

D. Empirical Evidence on the Internal and External Validity of SP Data Collected through the Cornell DCE

D.1. Validity Checks of the Quality of the Stated Preference Data

In sub-sections D.1 and D.2 we present empirical evidence on the internal and external validity of the SP data we collected through our DCE. Like other experimental research designs, the randomly assigned variation in product attributes in DCEs provides an internally valid research design to estimate the causal treatment effects of product attributes on subjects' stated preferences for tobacco products and quit attempts. However, because SP data are the subjects' responses about hypothetical choices, subjects might not provide thoughtful and meaningful responses that provide useful information about the actual choices they would make in real-world markets. In this section D.1, we report the results of validity checks on the quality of our SP data and the implications for the empirical results reported in the text of the paper.

As an overview of the sensitivity of the empirical results to the validity checks, Table D1 reports the sensitivity of conditional logit models of tobacco product choices to alternative approaches to improve SP data validity. Column (0) reports a baseline conditional logit model estimated over the same sample used in estimation of the main text models. Columns (1) – (3) report conditional logit models estimated using sample restrictions to improve the quality of the SP data. Column (4) reports a conditional logit model estimated using a combination of SP and revealed preference (RP) data. In columns (1) – (3) the point estimates of the alternative specific constants (ASCs) and the product attribute parameters tend to be very similar to the baseline model parameter estimates in column (0). The combined SP + RP model reported in column (4) yields estimated ASCs that are substantively different than the estimated ASCs in the baseline column (0) model. We will discuss the SP + RP data model results in more detail in the next sub-section D.2 of this Appendix.

In this sub-section we focus on sample restrictions that might improve the quality of our SP data. First, we examine data on the length of time subjects spent answering the DCE choice tasks, to identify possible "speedsters" who provided lower-quality responses. Figure D1 shows the distribution of time spent on the choice tasks. The median and mode times spent on the choice tasks are 2.9 and 1.9 minutes, respectively. Each subject completed 12 immediate choice tasks and 12 six-months-from-now choice tasks; the six-months-from-now choice tasks might be easier to complete quickly because they presented the subject with the same choices as in the preceding immediate choice task. For the model reported in column (1) of Table B1, we drop responses from 107 speedster subjects who completed the choice tasks in under 2 minutes.

Second, we examine data on the extent to which subjects paid attention to attribute variation across choice tasks. After subjects completed the choice tasks, we asked subjects which attributes varied across the tasks. Figure D2 shows the fraction of subjects who correctly indicated that the attribute in question varied. Although substantial fractions of the responses about attribute variation were incorrect, the results suggest that subjects paid the most attention to the price attribute of cigarettes, which was the most common tobacco product choice. The patterns of attentiveness across attributes and products are consistent with rational decisions to pay the most attention to the attributes and products that matter to their preferences. We also note that there is an ambiguity in our measure of attentiveness. As noted above, the six-months-from-now choice task was always identical to the preceding immediate choice task, i.e., in those pairs of tasks the

attributes did not vary. For the model reported in column (2) of Table D1, we drop responses from 148 subjects who were inattentive to variation in the attributes of cigarettes.

Third, we examine data on the extent to which subjects might have made irrational choices across the choice tasks. Lancsar and Louivere (2006) discuss conventional approaches to rationality in consumer demand theory and the application to SP data collected through DCEs. Our first validity check examines whether a subject's choices across different choice tasks violates the weak axiom of revealed preference (WARP), with respect to differences in the price attributes. An example of a WARP violation with respect to the price attribute would be if a subject chooses HTPs at a high price in one task, and in another task choose cigarettes even though the HTPs price is lower and the price of cigarettes is unchanged. Our second validity check examines WARP violations with respect to differences in the price attributes, the availability of flavors, and the availability of different levels of nicotine. An example of a WARP violation with respect to the price, flavor, and nicotine attributes would be if a subject chooses HTPs at a high price in one task, and in another task chooses cigarettes even though the HTPs price is lower, and other HTP attributes and the price of cigarettes are unchanged. Table D2 provides a summary of the criteria used to detect irrational choices and the fraction of irrational choices made by subjects in our DCE. Considering the 12 immediate choice tasks, the average subject made 1.71 choices that violated WARP with respect to prices, and 0.92 choices that violated WARP with respect to the price, flavor, and nicotine attributes. The number of choices that violated WARP with respect to prices is higher because some of those choices might reflect the subject's willingness to make tradeoffs between price and the other attributes. For the model reported in column (3) of Table D1, we drop responses from the 165 subjects who made any choices that violated WARP with respect to the price, flavor, and nicotine attributes.

We note that our SP data include subjects who were non-traders and chose the same tobacco product, usually cigarettes, in all choice tasks. Lancsar and Louivere (2006) argue that although the responses from non-traders do not help identify marginal rates of substitution between attributes, dropping non-traders from the sample might eliminate subjects with strong preferences for a product or attribute.

D.2 Improving Data Quality by Combining SP and RP Data

In this sub-section we discuss the approach reported in column (4) of Table D1, where we estimate a conditional logit model of tobacco product choices using a combination of SP and RP data. In his monograph on econometric analysis of discrete choice data, Train (2002, pp. 174-175) discusses the advantages and disadvantages of SP and RP data:

Revealed preference data have the advantage that they reflect actual choices.... However, RP data are limited to the choice situations and attributes of alternatives that currently exist or have existed historically. Often a researcher will want to examine people's responses in situations that do not currently exist, such as the demand for a new product. RP data are simply not available for these new situations.

Stated-preference data complement revealed-preference data.... The limitations of SP data are obvious: what people say they will do is often not the same as what they actually do. People might not know what they would do if a hypothetical situation were real. Or they might not be willing to say what they would do.

Train suggests that by combining RP and SP data, "the advantages of each can be obtained while mitigating the limitations. The SP data provide the needed variation in attributes, while the RP data ground the predicted shares in reality." He outlines the approach we take, where we use our DCE's subjects' responses about their tobacco product use and quit attempts over the past year as SP data to calibrate our model.

To construct RP choices, we propose a probabilistic classification rule that uses information from the background survey of respondents' smoking behaviors. The process and classification rule are described in table D3. The construction takes two steps, first, we classify respondents as either choosing cigarettes or HTPs. Second, we classify respondents as either choosing quit or not quit. To classify the choices of cigarettes and HTPs, we use information of subjects' vaping history, vaping status, and smoking status. Among all the subjects, those who have never vaped are classified as choosing cigarettes, among those who have ever vaped, if they currently do not vape at all, they are classified as choosing cigarettes, for those who vape ever day and someday, we classify their choices according to their smoking status. Specifically, among those who vape ever day, if they also smoke every day, then their probabilities of choosing cigarettes versus HTPs are 50% versus 50%, if they smoke someday (say m days out of 30 days), then their probabilities of choosing cigarettes versus HTPs are $m/(30 + m)$ versus $30/(30 + m)$. For those who vape someday, the probability rules are similar, and the choice of quitting is constructed following a similar rule, details are reported in table D3.

After obtaining RP choices, we jointly estimate the model with SP and RP data. We assume that the random error term for the RP data follows an i.i.d. type I extreme value distribution with scale parameter normalized to 1, and the scale parameter for the SP data is given by λ . The choice probabilities for individual i chooses alternative j in the RP and SP data can then be written as:

$$P_{ij}^{RP} = \frac{e^{(\alpha_j + X_{ij}^{RP} \beta)}}{\sum_k e^{(\alpha_k + X_{ik}^{RP} \beta)}}$$

$$P_{ij}^{SP} = \frac{e^{[\lambda(\alpha_j + X_{ij}^{SP} \beta)]}}{\sum_k e^{[\lambda(\alpha_k + X_{ik}^{SP} \beta)]}}$$

In the joint estimation we maximize the joint likelihood function:

$$L(\alpha, \beta, \lambda) = \sum_{i=1}^N \sum_{j=1}^J y_{ij}^{RP} \log P_{ij}^{RP} + \sum_{i=1}^N \sum_{j=1}^J y_{ij}^{SP} \log P_{ij}^{SP}$$

The results are presented in column (4) of table D1. The estimated scale parameter λ is 0.346, suggesting that the scales in the RP and SP data are somewhat different. However, the estimated coefficients on the product attributes show similar patterns as in the column (0) model based on SP data only. Overall, we interpret the results from the calibrated model as supporting the usefulness of our SP data.

Table D1. Estimation Results from Conditional Logit Models

		Immediate Choice Today				
		(0)	(1)	(2)	(3)	(4)
Alternative-Specific-Constant	Cigarettes	2.466*** (0.122)	2.447*** (0.136)	2.777*** (0.149)	2.888*** (0.161)	7.128*** (0.749)
Alternative-Specific-Constant	HTPs	2.049*** (0.131)	2.067*** (0.147)	2.222*** (0.154)	2.216*** (0.166)	5.961*** (0.533)
Price	Price in 100 JPY	-0.162***	-0.179***	-0.216***	-0.190***	-0.464***

		(0.010)	(0.012)	(0.013)	(0.013)	(0.052)
HTP available flavor	Tobacco, menthol, fruity/coffee/mint	0.092*	0.116**	0.187***	0.237***	0.196
		(0.051)	(0.059)	(0.063)	(0.059)	(0.308)
	Tobacco and menthol	0.016	0.022	0.081	0.114*	0.044
		(0.051)	(0.058)	(0.065)	(0.059)	(0.282)
HTP available nicotine level	Up to 30mg	0.004	-0.055	-0.005	0.012	0.017
		(0.049)	(0.056)	(0.063)	(0.054)	(0.152)
	Up to 50mg	-0.034	-0.060	-0.027	0.061	-0.153
		(0.052)	(0.058)	(0.066)	(0.057)	(0.175)
HTP warning	Are not completely risk free	0.016	0.015	0.023	0.010	0.045
		(0.063)	(0.071)	(0.075)	(0.073)	(0.181)
	Contain nicotine, which is addictive	-0.019	-0.010	-0.003	-0.097	-0.054
		(0.071)	(0.084)	(0.088)	(0.085)	(0.096)
	Smoking is a cause of lung cancer	0.027	0.056	0.081	-0.003	0.001
		(0.063)	(0.070)	(0.076)	(0.070)	(0.157)
Scale parameter of SP data						0.346***
						(0.034)
Log-likelihood at convergence		-5589	-4457	-3808	-3585	-5919
Number of subjects		523	416	375	358	523
Number of observations		18828	14976	13500	12888	18828

Notes: The reference category of HTP available flavor “tobacco only”, the reference category of HTP available nicotine level is “up to 10mg”, the reference category of HTP warning is “no warning”. Standard errors are clustered at the individual level. *** p<0.01, ** p<0.05, * p<0.1.

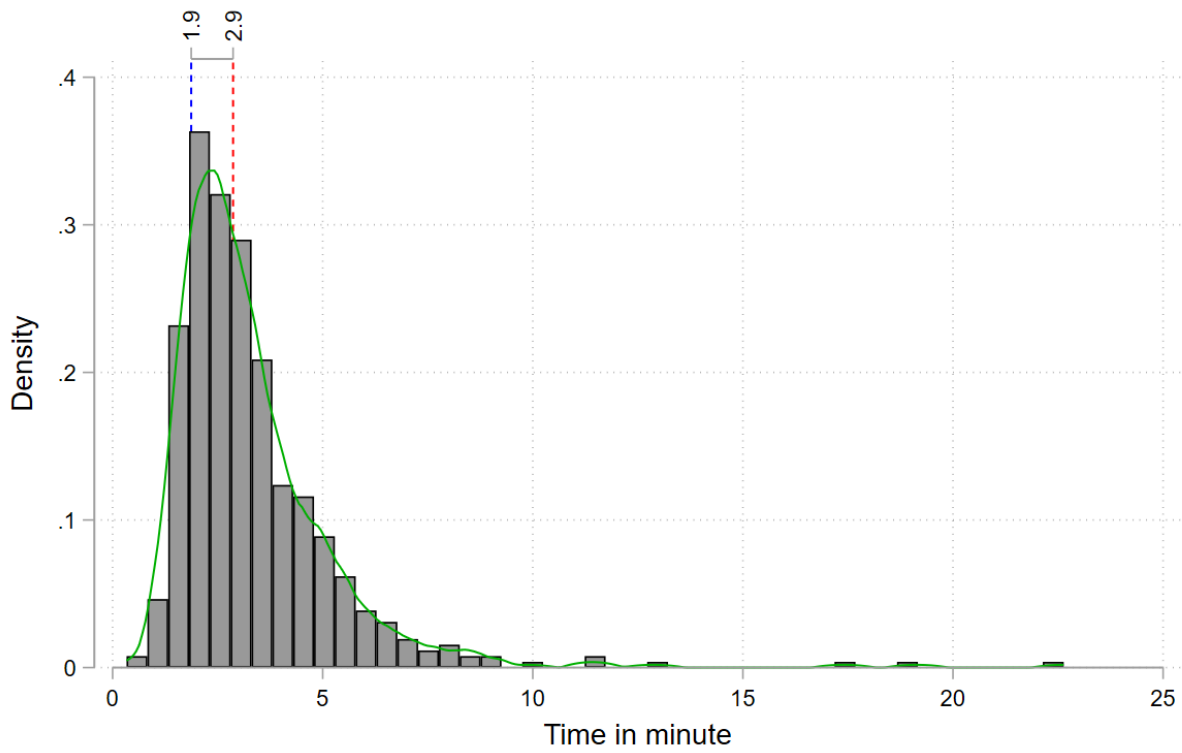


Figure D1. Distribution of Subject's Time Spent on the Choice Tasks

Notes: The blue and red dash line indicate the mode and median of time spent on the DCE section respectively. Out of 523 subjects, 1 is dropped due to zero measured time spent on the choice tasks, the histogram represents the distribution of 522 subject's time spent on the choice tasks.

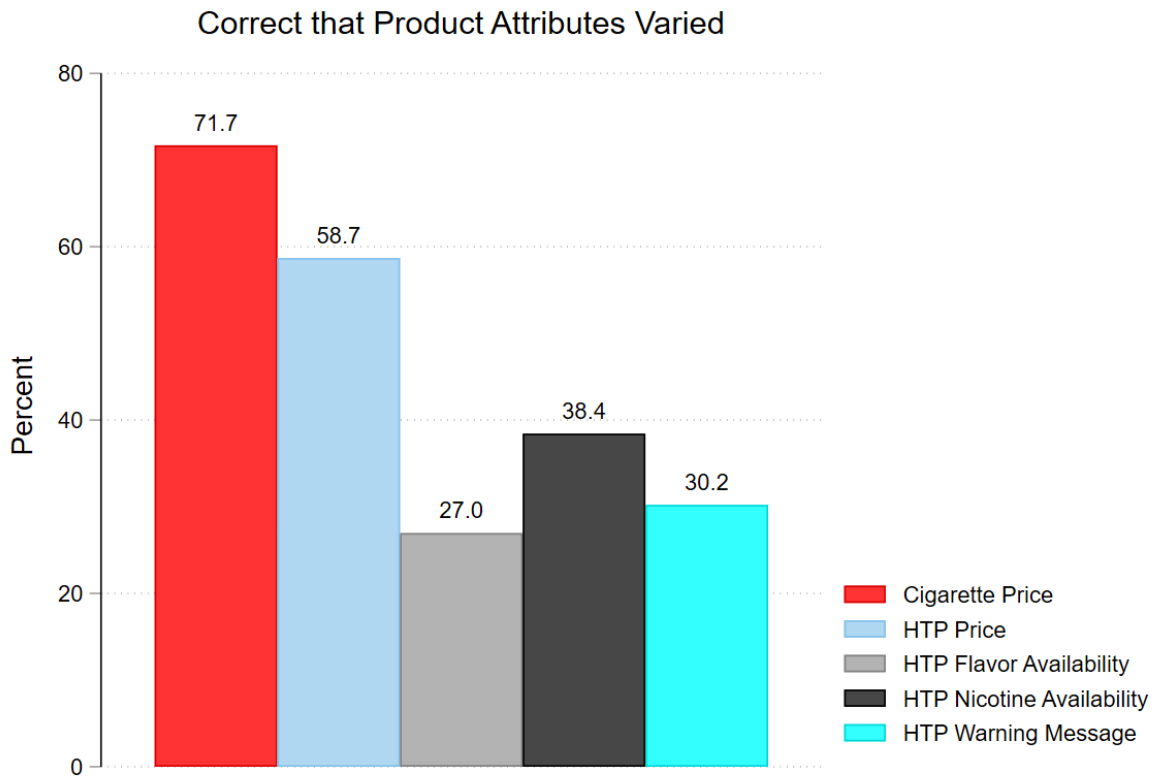


Figure D2. Fraction of Subjects that Were Correct that the Attribute in Question Varied.

Notes: Subjects are allowed to give correct but inconsistent answers, for example, they could select both price varied across scenarios and price was always the same. We use a strict criterion that the subject was correct that one attribute in question varied if she only selects the attribute varied and does not select the attribute was always the same.

Table D2. Criteria for Identifying Inconsistent Choices in the DCE

<u>Choices at any two tasks</u>		<u>Current vs previous attributes¹</u>		<u>Inconsistent?</u>	<u>% or # Inconsistent</u>		
at previous task	at current task	Cigarettes ²	HTPs ³			All attributes ³	Prices only
					Task level: % (<i>N</i> = 6,622)		
Cigarettes	HTPs	Same or better	Same or worse	Yes		2	4
Cigarettes	Quit	Better	(Any ⁴)	Yes		2	2
Cigarettes	Cigarettes	(Any)	(Any)	No			
HTPs	Cigarettes	Same or worse	Same or better	Yes		2	4
HTPs	Quit	(Any)	Better	Yes		0	1
HTPs	HTPs	(Any)	(Any)	No			
Quit	Cigarettes	Worse	(Any)	Yes		2	2
Quit	HTPs	(Any)	Worse	Yes		0	0
Quit	Quit	(Any)	(Any)	No			
					<i>Across all (6) rules</i>	8	14
					Subject level: # per subject (<i>N</i> = 602)	0.92 (range:0~10)	1.71 (range:0~11)

Notes:

¹ Product attributes at the current task compared to the previous task.

² Cigarettes' attribute includes price only; Cigarettes' attribute becoming "better (worse)" means Cigarette price decreases (increases) at the current task compared to the previous task.

³ HTPs' attributes include price, flavor availability, and nicotine levels; HTPs' attributes becoming "better (worse)" means HTP price decreases (increases), flavor availability increases (decreases), and nicotine level increases (decreases) at the current task compared to the previous task. Warnings of HTPs are not included in the criteria because subjects don't seem to have clear preference toward different warnings in the DCE.

⁴ Product attributes at the current task can be the same, better, or worse compared to the previous task.

Table D3. A Probabilistic Classification Rule to Construct Revealed Preference Choices

Step 1: classification of choices between cigarettes and e-cigarettes					
Subject	Vaping history	Vaping status	Smoking status	Probability of choosing cigarettes	Probability of choosing HTPs
N=523	Vape ever (Yes) N=378	Vape every day (N=216)	Smoke every day (N=188) Smoke someday (<i>m</i> out of 30 days) (N=28)	50% $m/(30 + m)$	50% $30/(30 + m)$
		Vape someday (<i>n</i> out of 30 days) (N=118)	Smoke every day (N=93) Smoke someday (<i>m</i> out of 30 days) (N=25)	$30/(30 + n)$ $m/(m + n)$	$n/(30 + n)$ $n/(m + n)$
		Not at all (N=44)	Smoke every day (N=40) Smoke someday (N=4)	100%	0
	Vape every (No) N=145	Smoke every day (N=122) Smoke someday (N=23)			
Step 2: classification of choices of quitting					
Subject	Quitting intention	Quitting history	Probability of choosing quitting		
N=523	Plan to quit (Yes) N=149	Tried quitting (N=107) Didn't try to quit (N=42)	1/26		
	Plan to quit (No) N=374	Tried quitting (N=30) Didn't try to quit (N=344)	1/52 0		

Appendix E: Age Cohort Descriptive Tables

There are notable differences present across the two age cohorts in the descriptive tables reported in Appendix E. Table E1 presents DCE choices for the entire sample and separately for each age group. Table E2 reports demographic characteristics by cohort. The younger cohort has an average age of 36, contains more women 32.3 percent, has a larger household size 3.2, is more likely to be working fulltime 77 percent, and is more likely to earn more than 5000k yen 71 percent. Comparatively, the older cohort has an average age of 63, only 18.3 percent are women, the average household size is 2.6, only 45.6 percent are full time employed, and 52.4 percent earn more than 5000k per year. The lower rates of full-time employment and lower earnings suggest a sizable portion of our older cohort is retired.

Consistent with Levy et al. (2024), the older cohort reports smoking characteristics (Table E3) that depict higher rates of ‘smoking prevalence’. Smokers in the older cohort are more likely to be everyday smokers (86.3 percent), less likely to have tried to quit smoking in the last year (24.1 percent), and less likely to report an intention to quit smoking in the next 6 months (22.8 percent). Comparatively, smokers in the younger cohort are slightly less likely to be everyday smokers (83.3 percent), slightly more likely to have tried to quit smoking in the last year (32.3), and slightly more likely to report an intention to quit smoking in the next 6 months (29.1 percent). In both groups, a majority of respondents are everyday smokers who are unlikely to have attempted to quit smoking in the past year and are unlikely to attempt to quit smoking in the next six months. However, the older cohort is consistently more likely to report behaviors associated with higher ‘smoking prevalence’.

While not the same age cohorts reported in Levy et al. (2024), we observe similar trends, where the younger cohort is far more likely to have ever used an HTP and is more likely to use HTPs daily. As reported in Table E4, among respondents aged 20-49, 83.3 percent have ever used an HTP, compared to 59.3 percent of respondents over the age of 50. Among respondents who have used an HTP, 63.8 percent of respondents aged 20-49 consider themselves to be an everyday HTP users, while only 46.2 percent of respondents over the age of 50 did.

In terms of their responses to the choice scenarios (Table E5), the younger cohort is almost equally likely to have chosen cigarettes today, 46.6 percent of the time, as they are to have chosen HTP today, 46.3 percent of the time. Their six months from now responses show a slight increase in willingness to choose HTPs, 47.6 percent of the time, compared to cigarettes, 44.9 percent of the time. The older cohort is far more likely to choose cigarettes now and in their choice six months from now. Among respondents over the age of 50, 53.6 percent of the time their immediate choice was cigarettes, compared to an immediate choice of HTPs, 30.5 percent of the time. In six months from now responses, the older cohort was only slightly more likely to choose HTPs, 31.1 percent of the time, compared to the choice of cigarettes, 52 percent of the time.

Table E1: Descriptive Statistics of DCE Attributes by Age Cohort

	All	Ages 20-49	Ages 50 plus
Cigarette price in Yen	615.46 (350.78)	614.74 (359.44)	616.30 (340.44)
<i>HTP price levels</i>			
HTP price = 235 yen	0.336 (0.47)	0.330 (0.47)	0.342 (0.47)
HTP price = 470 yen	0.339 (0.47)	0.334 (0.47)	0.345 (0.48)
HTP price = 940 yen	0.325 (0.47)	0.336 (0.47)	0.313 (0.46)
<i>HTP health messages</i>			
No warning	0.251 (0.43)	0.246 (0.43)	0.257 (0.44)
Small fraction of risks of cigarettes	0.255 (0.44)	0.253 (0.43)	0.257 (0.44)
Contains nicotine, which is addictive	0.243 (0.43)	0.248 (0.43)	0.236 (0.42)
Smoking is a cause for lung cancer	0.252 (0.43)	0.252 (0.43)	0.251 (0.43)
<i>HTP nicotine content levels</i>			
Nicotine content up to 10mg	0.333 (0.47)	0.333 (0.47)	0.333 (0.47)
Nicotine content up to 30mg	0.333 (0.47)	0.333 (0.47)	0.333 (0.47)
Nicotine content up to 50mg	0.333 (0.47)	0.333 (0.47)	0.333 (0.47)

HTP flavor levels

Tobacco, menthol, mentholated fruity	0.333 (0.47)	0.333 (0.47)	0.333 (0.47)
Tobacco and menthol	0.333 (0.47)	0.333 (0.47)	0.333 (0.47)
Tobacco only	0.333 (0.47)	0.333 (0.47)	0.333 (0.47)
N	6,276	3,384	2,892

Notes: Data from Japan Discrete Choice Experiment (2021). Reported statistics include Mean (Standard Deviation).

Table E2: Descriptive Statistics of Demographics by Age Cohort

	All	Ages 20-49	Ages 50 plus
Female	0.258 (0.44)	0.323 (0.47)	0.183 (0.39)
Age	48.535 (15.13)	36.287 (7.59)	62.867 (6.97)
Household size	2.906 (1.39)	3.202 (1.47)	2.560 (1.19)
<i>Education</i>			
Junior high school	0.023 (0.15)	0.032 (0.18)	0.012 (0.11)
High school	0.254 (0.44)	0.245 (0.43)	0.266 (0.44)
Vocational school	0.092 (0.29)	0.085 (0.28)	0.100 (0.30)
Junior college	0.031 (0.17)	0.018 (0.13)	0.046 (0.21)
Some undergrad	0.021 (0.14)	0.011 (0.10)	0.033 (0.18)
Undergraduate	0.541 (0.50)	0.564 (0.50)	0.515 (0.50)
Postgraduate	0.038 (0.19)	0.046 (0.21)	0.029 (0.17)
Full time employed	0.625 (0.48)	0.770 (0.42)	0.456 (0.50)
Income above 5000k yen	0.626 (0.48)	0.710 (0.45)	0.524 (0.50)
N	523	282	241

Notes: Data from Japan Discrete Choice Experiment (2021). Reported statistics include Mean (Standard Deviation).

Table E3: Descriptive Statistics of Smoking History by Age Cohort

	All	Ages 20-49	Ages 50 plus
Everyday smoker	0.847 (0.36)	0.833 (0.37)	0.863 (0.34)
Age of smoking initiation	23.21 (9.01)	22.035 (5.01)	24.585 (11.97)
<i>On avg, how many cigs per day?</i>			
0-5	0.212 (0.41)	0.238 (0.43)	0.183 (0.39)
6-10	0.289 (0.45)	0.312 (0.46)	0.261 (0.44)
11-15	0.218 (0.41)	0.209 (0.41)	0.228 (0.42)
16-20	0.191 (0.39)	0.138 (0.35)	0.253 (0.43)
21-25	0.044 (0.21)	0.06 (0.24)	0.025 (0.16)
26-30	0.023 (0.15)	0.014 (0.12)	0.033 (0.18)
31-35	0.008 (0.09)	0.011 (0.1)	0.004 (0.06)
36-40	0.01 (0.1)	0.011 (0.1)	0.008 (0.09)
Above 40	0.006 (0.08)	0.007 (0.08)	0.004 (0.06)
<i>How soon do you smoke after waking up</i>			
Within 5 minutes	0.189 (0.39)	0.174 (0.38)	0.207 (0.41)
6-30 minutes	0.436 (0.5)	0.415 (0.49)	0.461 (0.5)
31-60 minutes	0.149 (0.36)	0.163 (0.37)	0.133 (0.34)
1-2 hours	0.117 (0.32)	0.11 (0.31)	0.124 (0.33)
2-3 hours	0.044 (0.21)	0.053 (0.22)	0.033 (0.18)
3-4 hours	0.015 (0.12)	0.018 (0.13)	0.012 (0.11)
More than 4 hours	0.05 (0.22)	0.067 (0.25)	0.029 (0.17)
<i>Usual smoking flavor</i>			
Menthol	0.369 (0.48)	0.454 (0.5)	0.27 (0.44)

Non-menthol	0.512 (0.5)	0.457 (0.5)	0.577 (0.49)
Other flavor	0.008 (0.09)	0.007 (0.08)	0.008 (0.09)
No usual type	0.111 (0.31)	0.082 (0.27)	0.145 (0.35)
Did not try quitting in past 12 months	0.715 (0.45)	0.677 (0.47)	0.759 (0.43)
Not planning to quit in next 6 months	0.738 (0.44)	0.709 (0.45)	0.772 (0.42)
<hr/> N	<hr/> 523	<hr/> 282	<hr/> 241

Notes: Data from Japan Discrete Choice Experiment (2021). Reported statistics include Mean (Standard Deviation).

Table E4: Descriptive Statistics of HTP Use by Age Cohort

	All	Ages 20-49	Ages 50 plus
Have you ever used an HTP	0.723 (0.45) [523]	0.833 (0.37) [282]	0.593 (0.49) [241]
Everyday HTP user	0.571 (0.49) [378]	0.638 (0.48) [235]	0.462 (0.50) [143]
Uses any flavored HTP	0.683 (0.47) [334]	0.717 (0.45) [223]	0.613 (0.49) [111]
Use tobacco flavored HTP	0.527 (0.50) [334]	0.556 (0.50) [223]	0.468 (0.50) [111]
Use menthol flavored HTP	0.575 (0.49) [334]	0.610 (0.49) [223]	0.505 (0.50) [111]
Use mentholated fruity flavored HTP	0.225 (0.42) [334]	0.265 (0.44) [223]	0.144 (0.35) [111]
Use mint flavored HTP	0.120 (0.32) [334]	0.135 (0.34) [223]	0.090 (0.29) [111]
Use coffee flavored HTP	0.060 (0.24) [334]	0.058 (0.23) [223]	0.063 (0.24) [111]

Notes: Data from Japan Discrete Choice Experiment (2021). Reported statistics include Mean (Standard Deviation) [N].

Table E5: Descriptive Statistics of DCE Choice Now and in Six months by Age Cohort

	All	Ages 20-49	Ages 50 plus
Choose cig now	0.498 (0.50)	0.466 (0.50)	0.536 (0.50)
Choose HTP now	0.390 (0.49)	0.463 (0.50)	0.305 (0.46)
Quit now	0.112 (0.31)	0.071 (0.26)	0.159 (0.37)
Choose cig in 6 months	0.482 (0.50)	0.449 (0.50)	0.520 (0.50)
Choose HTP in 6 months	0.400 (0.49)	0.476 (0.50)	0.311 (0.46)
Choose quit in 6 months	0.118 (0.32)	0.075 (0.26)	0.169 (0.37)
N	6,276	3,384	2,892

Notes: Data from Japan Discrete Choice Experiment (2021). Reported statistics include Mean (Standard Deviation).

Table E6: Descriptive Statistics of Post DCE Perceptions of Smoking and HTP use by Age Cohort

	All	Ages 20-49	Ages 50 plus
Perception: Life lost smoking	7.164 (7.86)	9.270 (8.56)	4.701 (6.08)
Perception: Life lost using HTPs	6.430 (6.92)	7.770 (7.61)	4.863 (5.61)
<i>Comparing to cigs, HTPs are</i>			
Much less harmful	0.140 (0.35)	0.142 (0.35)	0.137 (0.34)
Less harmful	0.503 (0.50)	0.539 (0.50)	0.461 (0.50)
Just as harmful	0.212 (0.41)	0.234 (0.42)	0.187 (0.39)
More harmful	0.019 (0.14)	0.025 (0.16)	0.012 (0.11)
Much more harmful	0.013 (0.11)	0.007 (0.08)	0.021 (0.14)
I don't know	0.113 (0.32)	0.053 (0.22)	0.183 (0.39)
<i>I smoke more than I should</i>			
Strongly agree	0.055 (0.23)	0.067 (0.25)	0.041 (0.20)
Somewhat agree	0.283 (0.45)	0.323 (0.47)	0.237 (0.43)
Neither disagree nor agree	0.293 (0.45)	0.287 (0.45)	0.299 (0.46)
Somewhat disagree	0.277	0.262	0.295

	(0.45)	(0.44)	(0.46)
Strongly disagree	0.090	0.060	0.124
	(0.29)	(0.24)	(0.33)
I don't know	0.002	0.000	0.004
	(0.04)	(0.00)	(0.06)
<i>Compared to tobacco & menthol flavors, using other HTP flavors are</i>			
Much less harmful	0.094	0.092	0.095
	(0.29)	(0.29)	(0.29)
Less harmful	0.390	0.433	0.340
	(0.49)	(0.50)	(0.47)
Just as harmful	0.338	0.376	0.295
	(0.47)	(0.48)	(0.46)
More harmful	0.021	0.018	0.025
	(0.14)	(0.13)	(0.16)
Much more harmful	0.011	0.004	0.021
	(0.11)	(0.06)	(0.14)
I don't know	0.145	0.078	0.224
	(0.35)	(0.27)	(0.42)
<hr/>			
N	523	282	241

Notes: Data from Japan Discrete Choice Experiment (2021). Reported statistics include Mean (Standard Deviation).

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